

The Factor-copula Framework (1/3)

The random variables $x_{n,1 \leq n \leq N}$ represent the creditworthiness of the obligors in a reference portfolio of N risky obligors, and x_n 's are correlated via the common factors $\mathbf{Z} = [Z_1, Z_2, \dots, Z_d]^d$, i.e. for $n = 1, \dots, N$:

$$x_n = \beta_n^T \mathbf{Z} + b_n \varepsilon_n, \quad (1)$$

where

- ▶ $\beta_n = [\beta_{n,1}, \beta_{n,2}, \dots, \beta_{n,d}]^T$ are the weighting coefficients of the common factors \mathbf{Z} ;
- ▶ $b_n = \sqrt{1 - \sum_{i=1}^d \beta_{n,i}^2}$ are the weighting coefficients of the idiosyncratic factor ε_n ;
- ▶ The common factors are assumed to be independent to the idiosyncratic factor, i.e. $\mathbf{Z} \perp \varepsilon_n$.

The Factor-copula Framework (2/3)

$$x_n = \beta_n^T \mathbf{Z} + b_n \varepsilon_n,$$

- ▶ The obligor n defaults if and only if x_n is less than the default threshold, given as the inverse of its CDF;
- ▶ The common factors \mathbf{Z} can be independent to each other or governed by a copula, which can impose a different marginal distribution for than that of ε_n ;
- ▶ In the old literature, Gaussian copula is usually used, i.e. $\mathbf{Z}, \varepsilon_n \sim N(0, 1)$;
- ▶ In recent years, Student-t copula has been adopted in the industry at least for benchmark purpose, i.e. $\mathbf{Z}, \varepsilon_n \sim F_t(0, 1)$;

The Factor-copula Framework (3/3)

- ▶ There are factor-copula models where the systematic factors and the idiosyncratic factors follow different distributions, e.g.
 - Oh and Patten 2017
 - 2019 ECB Guide to Internal Models
- ▶ To illustrate the generic method, in our example we assume the systematic factors follow a d -variate t distribution. The idiosyncratic factors remain Gaussian. (It can be made the other way around or assumed to follow other distributions)
 - A more convenient formulation of the hybrid copula model:

$$x_n = \sqrt{W} \beta_n^T \mathbf{Z} + b_n \varepsilon_n. \quad (2)$$

where W has an inverse gamma distribution, i.e., $W \sim I_g(\nu/2, \nu/2)$. ν is the degrees of freedom.

Mathematical Formulation of the Problem-to-solve

The portfolio loss L is defined as

$$L = \sum_{n=1}^N \mathbf{1}_{x_n \leq \xi_n} \cdot l_n \quad (3)$$

Then the problem-to-solve comprise two parts

- ▶ Risk quantification: Value-at-Risk (VaR) and ES of L .
- ▶ Risk allocation: Euler risk allocation of VaR and ES.
 - A risk measure is decomposed as the sum of risk contributions of the obligors/sub-portfolios in the reference portfolio.
 - Homogenous: scaling the risk measure by a constant changes the risk decomposition by the same scale. E.g., increasing the loss-at-default of all the obligors by 10% would increase VaR/ES by 10% and the risk contribution of a certain obligor should also increase by 10%.

Recall the COS Method (1/2)

The essence of the COS method is that, a probability density function can be recovered from a truncated Fourier cosine series, of which the coefficients can be extracted from the characteristics function (ch.f), and thus, are readily available.

- ▶ That is, within the truncation range $[a, b]$ of a density function f , we have

$$f(x) \approx \sum'_{k=0}^K A_k \cos\left(k\pi \frac{x-a}{b-a}\right), \quad (4)$$

where

$$A_k = \frac{2}{b-a} \operatorname{Re} \left\{ \varphi\left(\frac{k\pi}{b-a}\right) \cdot \exp\left(-i \frac{ka\pi}{b-a}\right) \right\}$$

with $\varphi(\cdot)$ being the ch.f. of $f(x)$, and \sum' indicates that the first term in the sum is weighted by one-half.

Recall the COS Method (2/2)

To apply COS to the portfolio loss distribution of a multifactor copula model, we

1. first numerically evaluate the ch.f at a grid of points in the Fourier domain, i.e., $k\pi \frac{x-a}{b-a}$, $0 \leq k \leq K$.
2. then reconstruct the CDF function of the loss by COS, i.e.

$$F(y) = \int_a^y f(x)dx = \frac{A_0}{2}(y-a) + \sum_{k=1}^K A_k \frac{b-a}{k\pi} \sin\left(k\pi \frac{y-a}{b-a}\right) \quad (5)$$

It does not rely on the assumption of Gaussian distributions!

The COS Approach for Risk Quantification

The key is to solve the characteristic function (ch.f.) of the portfolio loss L . The ch.f. is derived as follows:

1. Conditional on the common factors, defaults of the obligors are independent Bernoulli random variables. Thus the conditional ch.f of the total loss $L = \sum l_n \mathbf{1}_{x_n \leq \xi_n}$ is

$$\mathbb{E}[\varphi_L(\omega)|Z] = \prod_{n=1}^N \mathbb{E}\left[e^{i\omega l_n \cdot \mathbf{1}_{\varepsilon_n \leq \alpha_n(z_n)}}\right] \quad (6)$$

where $\alpha_n(\mathbf{z}_n) = \frac{\xi_n - \beta_n^T \mathbf{z}}{b_n}$

2. Each expectation $\mathbb{E}\left[e^{i\omega l_n \cdot \mathbf{1}_{\varepsilon_n \leq \alpha_n(z_n)}}\right]$ in the product is simply given by the analytical expression of the Bernoulli ch.f.
3. Finally, the ch.f of the portfolio loss distribution can be obtained from the conditional ch.f $\mathbb{E}[\varphi_L(\omega)|Z]$ by numerically integrating out the common factors \mathbf{Z} .

Risk Measures via the COS-recovered CDF

- ▶ VaR : Very simple! Given the recovered CDF of the portfolio loss, the q -th quantile can be solved numerically, e.g., solving $P(L \leq \theta) = q$ via a root-searching algorithm
- ▶ ES: An analytical expression for ES is available, by integrating the loss with respect to the Fourier series expansion of CDF.

Euler Risk Allocation of ES

Conditional ES decomposes the ES by the Euler principle for risk allocation. We consider the following definition:

$$\text{CES}_n = \mathbb{E}[\mathbf{1}_{x_n \leq \xi_n} \cdot l_n \mid L \geq \text{VaR}_\alpha].$$

Such that

$$\text{ES} = \mathbb{E} \left[\sum_n \mathbf{1}_{x_n \leq \xi_n} \cdot l_n \mid L \geq \text{VaR}_\alpha \right] = \sum_n \text{CES}_n$$

Our Solution for Euler Risk Allocation of ES (1/2)

Apply Bayes law, we yield that

$$\begin{aligned} \text{CES}_n &= I_n \cdot P(x_n \leq \xi_n | L \geq \text{VaR}_\alpha) \\ &= I_n \cdot \frac{P(x_n \leq \xi_n, L \geq \text{VaR}_\alpha)}{P(L \geq \text{VaR}_\alpha)} \\ &= \frac{I_n \cdot p_n}{\alpha} \cdot P(L \geq \text{VaR}_\alpha | x_n \leq \xi_n). \end{aligned} \quad (7)$$

$P(L \geq \text{VaR}_\alpha | x_n \leq \xi_n)$ can be solved by the COS approach again as for VaR and ES!

Our Solution for Euler Risk Allocation of ES (2/2)

The coding of the COS calculation for $P(L \geq \text{VaR}_\alpha | x_n \leq \xi_n)$ can be easily integrated into the coding of the COS calculation for the portfolio loss distribution.

$$\begin{aligned}\varphi_{n,L}(\omega) &= \mathbb{E} \left[e^{i\omega L} \mid x_n \leq \xi_n \right] \\ &= \frac{\mathbb{E} \left[e^{i\omega L} \cdot \mathbf{1}_{x_n \leq \xi_n} \right]}{P(x_n \leq \xi_n)} \\ &= \frac{1}{p_n} \mathbb{E} \left[\mathbb{E} \left[e^{i\omega L} \cdot \mathbf{1}_{x_n \leq \xi_n} \mid \mathbf{Z} = \mathbf{z} \right] \right] \\ &= \frac{1}{p_n} \mathbb{E} \left[\left(\prod_{j \neq n} \mathbb{E} \left[e^{i\omega l_j \cdot \mathbf{1}_{\varepsilon_j \leq \alpha_j(z_j)}} \mid \mathbf{Z} = \mathbf{z} \right] \right) \right. \\ &\quad \left. \cdot \mathbb{E} \left[e^{i\omega l_n \cdot \mathbf{1}_{\varepsilon_n \leq \alpha_n(z_n)}} \cdot \mathbf{1}_{\varepsilon_n \leq \alpha_n(z_n)} \mid \mathbf{Z} = \mathbf{z} \right] \right] \quad (8)\end{aligned}$$

Euler Risk Allocation

Conditional VaR decomposes the VaR by the Euler principle for risk allocation.

$$\text{CVaR}_n = \mathbb{E}[\mathbf{1}_{x_n \leq \xi_n} \cdot I_n \mid L = \text{VaR}_\alpha].$$

Such that

$$\text{VaR} = \mathbb{E} \left[\sum_n \mathbf{1}_{x_n \leq \xi_n} \cdot I_n \mid L = \text{VaR}_\alpha \right] = \sum_n \text{CVaR}_n$$

The rest of the calculation follows the same steps as for the Euler risk allocation of ES, except that we need to evaluate the expectation conditional on a small neighborhood around VaR_α .

Our Solution for Conditional VaR

Similar to Conditional ES, Conditional VaR decomposes the VaR by the Euler principle for risk allocation. We consider the following similar definition:

$$\text{CVaR}_n = \mathbb{E}[\mathbf{1}_{x_n \leq \xi_n} \cdot l_n | L = \text{VaR}_\alpha].$$

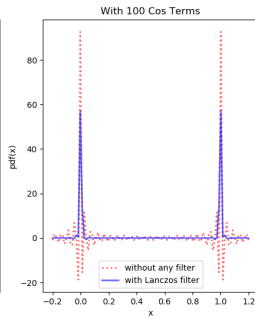
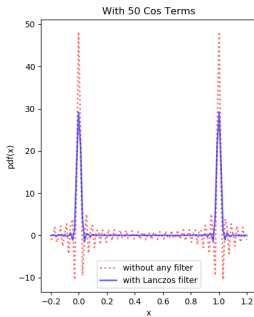
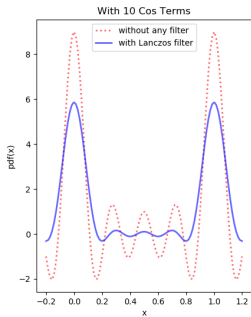
So that

$$\text{VaR} = \mathbb{E}\left[\sum_n \mathbf{1}_{x_n \leq \xi_n} \cdot l_n \middle| L = \text{VaR}_\alpha\right] = \sum_n \text{CVaR}_n$$

It then follows that

$$\begin{aligned}\text{CVaR}_n &= l_n \cdot P(x_n \leq \xi_n | L = \text{VaR}_\alpha) \\ &= l_n \cdot \frac{P(x_n \leq \xi_n, L = \text{VaR}_\alpha)}{P(L = \text{VaR}_\alpha)} \\ &\approx l_n \cdot \frac{P(x_n \leq \xi_n, \text{VaR}_\alpha - \epsilon \leq L \leq \text{VaR}_\alpha + \epsilon)}{P(\text{VaR}_\alpha - \epsilon \leq L \leq \text{VaR}_\alpha + \epsilon)} \\ &= l_n \cdot p_n \cdot \frac{P(\text{VaR}_\alpha - \epsilon \leq L \leq \text{VaR}_\alpha + \epsilon | x_n \leq \xi_n)}{P(\text{VaR}_\alpha - \epsilon \leq L \leq \text{VaR}_\alpha + \epsilon)}\end{aligned}\tag{9}$$

Example of Bernoulli Distribution



Error Analysis - 1/7

- ▶ Denote the possible realizations of L by $\{0 \leq L_0 \leq L_1, \dots, L_m, \dots, \leq L_M \leq \pi\}$.
- ▶ Applying the COS expansion to have

$$f_L(x) = \sum_{k=0}^{\infty} A_k \cos kx$$

with

$$\begin{aligned} A_k &= \frac{2}{\pi} \operatorname{Re} \{ \varphi(k) \} = \frac{2}{\pi} \operatorname{Re} \left\{ \sum_{m=0}^M e^{ikL_m} p_m \right\} \\ &= \frac{2}{\pi} \sum_{m=0}^M \cos(kL_m) p_m \end{aligned} \quad (11)$$

where p_m is the probability of $L = L_m$.

Error Analysis - 2/7

- ▶ Thus the Fourier cosine expansion of the probability density of L is

$$\begin{aligned} f_L(x) &= \sum_{k=0}^{\infty} \frac{2}{\pi} \sum_{m=0}^M \cos(kL_m) p_m \cos kx \\ &= \sum_{m=0}^M p_m \sum_{k=0}^{\infty} \frac{2}{\pi} \cos(kL_m) \cos kx \\ &= \sum_{m=0}^M p_m f_m(x) \end{aligned} \tag{12}$$

where

$$f_m(x) = \sum_{k=0}^{\infty} \frac{2}{\pi} \cos(kL_m) \cos kx$$

Error Analysis - 3/7

- ▶ Integrating f_L from 0 gives the COS CDF of L

$$\begin{aligned} F_L(x) &= \sum_{m=0}^M p_m \left[\frac{1}{\pi} x + \sum_{k=1}^{\infty} \frac{2}{k\pi} \cos(kL_m) \sin kx \right] \\ &= \sum_{m=0}^M p_m F_m(x) \end{aligned} \quad (13)$$

where

$$F_m(x) = \frac{1}{\pi} x + \sum_{k=1}^{\infty} \frac{2}{k\pi} \cos(kL_m) \sin kx$$

Error Analysis - 4/7

Then we truncate the number of series term to K and modify the series coefficients by the filter to have

$$f_L^\sigma(x) = \sum_{m=0}^M p_m f_m^\sigma(x) \quad (14)$$

and

$$F_L^\sigma(x) = \sum_{m=0}^M p_m F_m^\sigma(x) \quad (15)$$

where

$$f_m^\sigma(x) = \sum_{k=0}^{K-1} \frac{2}{\pi} \sigma(k/K) \cos(kL_m) \cos kx$$

and

$$F_m^\sigma(x) = \frac{1}{\pi} x + \sum_{k=1}^K \frac{2}{k\pi} \sigma(k/K) \cos(kL_m) \sin kx$$

Error Analysis - 4/7

The key insight here is that on $[-\pi, \pi]$,

$$F_0(x) = \frac{1}{\pi}x + \sum_{k=1}^{\infty} \frac{2}{k\pi} \sin kx$$

is the Fourier series expansion of the function

$$H_0(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq \pi \\ -1 & \text{if } -\pi \leq x < 0 \end{cases}$$

and that

$$F_m(x) = \frac{1}{\pi}x + \sum_{k=1}^{\infty} \frac{2}{k\pi} \cos(kL_m) \sin kx$$

is the Fourier series expansion of the function

$$H_m(x) = \begin{cases} 1 & \text{if } L_m \leq x \leq \pi \\ 0 & \text{if } -L_m < x < L_m \\ -1 & \text{if } -\pi \leq x < -L_m \end{cases}$$

Error Analysis - 6/7

The convergence speed of the Fourier series expansion with spectral filter for a piecewise constant function is governed by the convergence order of the filter, as proven in [Vandeven 1991]:

- ▶ If we have a function $f \notin C^{p-1}$, ie, if $f(y)$ has a jump discontinuity at one or more points of order smaller than, or equal to, $p - 1$, the following estimate holds:

$$f_N^\sigma(y) - f(y) \sim O\left(N^{1/2-p}\right).$$

Given that the CDF of L is a linear combination of H_m , weighted by p_m , it follows that F_L^σ converges to the true CDF of L at the speed as described above.

Error Analysis - 7/7

Recall that there is one extra layer of approximation: the cosine coefficients A_k is obtained via numerical integration based on Clenshaw-Curtis rule after we truncate the integration range with a truncation error at the level of TOL . Let us denote the error term from this numerical integration part as $\epsilon(N, TOL)$, which depends on the the number of integration points N and the range truncation tolerance TOL . Then it can be shown that this error term propogates in our approximation of the CDF as follows:

$$\hat{F}_m^\sigma(x) = F_m^\sigma(x) + O(K) \cdot \epsilon(N, TOL).$$

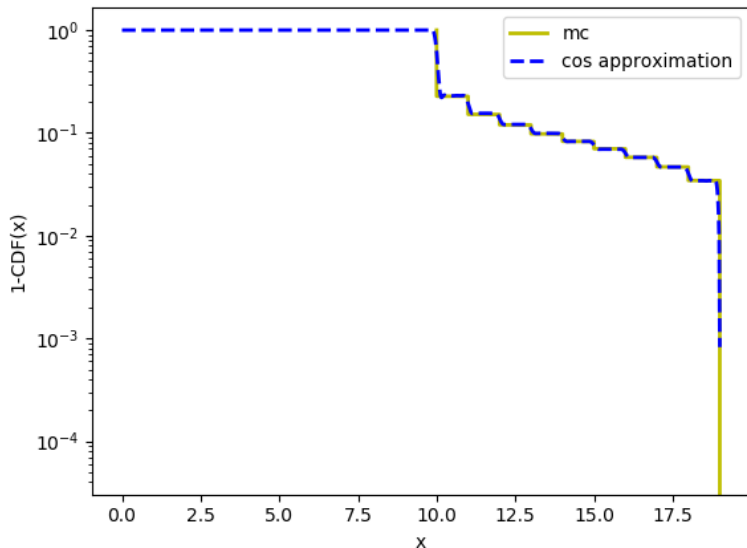
Numerical Example 1: A Small portfolio

To observe the behavior of the COS method, we first constructed a simple portfolio with 10 obligors, one of which creates name concentration. We consider a two-factor Gaussian copula, and a hybrid copula with Student-t distribution for the systematic factors and Gaussian distribution for the idiosyncratic factors.

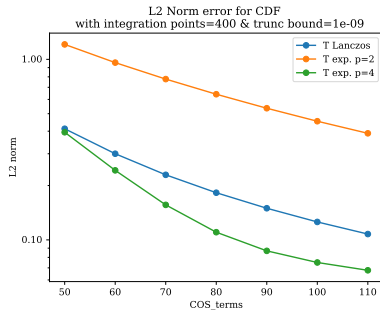
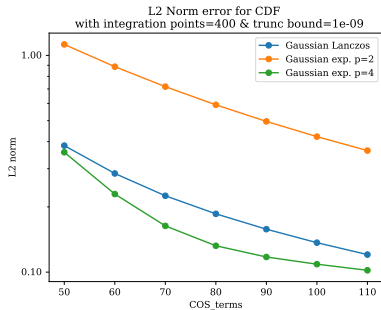
- ▶ Number of obligors: 10
- ▶ $\beta_{n,1} = 0.8, \beta_{n,2} = 0.4$
- ▶ $p_1 = 0.01, p_n = 0.001, n = 2, \dots, N$
- ▶ $l_1 = 10, l_n = 1, n = 2, \dots, N$
- ▶ Degree of freedom in the t Copula: 8.

CDF Conditional on Default of One Name

Conditional distribution under t with 200 COS terms and 100 int. pts

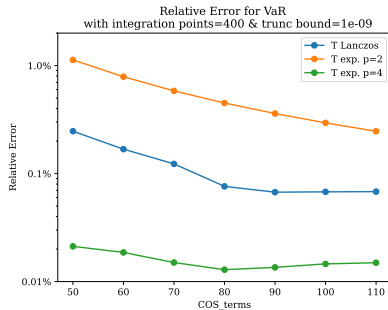
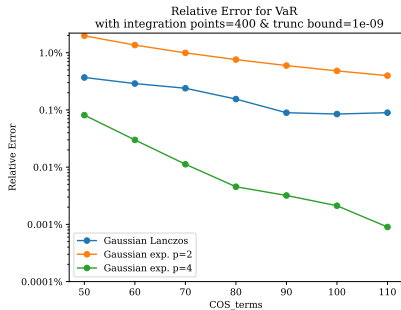


CDF of Portfolio Loss



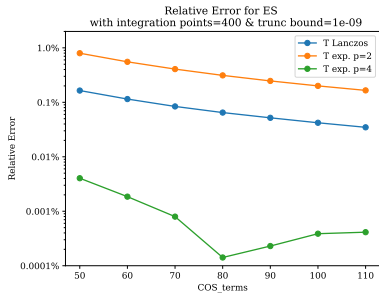
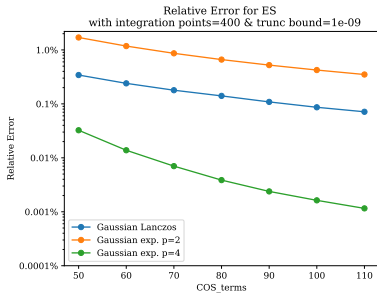
Left: Gaussian copula. Right: T copula.

VaR of Portfolio Loss



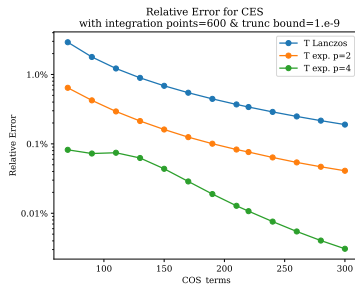
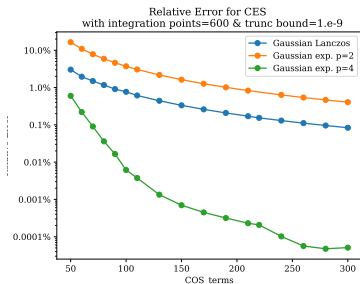
Left: Gaussian copula. Right: T copula.

ES of Portfolio Loss



Left: Gaussian copula. Right: T copula.

CES of Portfolio Loss



Left: Gaussian copula. Right: T copula.

Performance

Copula	Method	Time (sec.)	VaR Val.	ES Val.	Abs. Err. VaR	Abs. Err. ES	Rel. Err. VaR	Rel. Err. ES
Gaussian-t hybrid	COS (Nr. COS terms = 110, exp. filter p=4)	28	2,333.61	2,810.32	0.55	3.85	0.02%	0.14%
	MC	413	2,340.97	2,822.40	7.91	8.23	0.34%	0.29%
Gaussian	COS (Nr. COS terms = 110, exp. filter p=4)	20	1,967.53	2,215.72	0.16	0.79	0.01%	0.04%
	MC	268	1,971.86	2,224.60	4.49	9.67	0.23%	0.44%

Reference



Dominic O'kane

Modelling single-name and multi-name credit derivatives
The Wiley Finance Series, (2008).



G.C. Papiol, L.O.Gracia and C.W. Oosterlee

Quantifying credit portfolio losses under multifactor models
International J of Computer Mathematics, 96(11): 2135-2156, 2019. <https://doi.org/10.1080/00207160.2018.1447666>.



F. Fang and C.W. Oosterlee

A Novel Pricing Method for European Options Based on
Fourier-Cosine Series Expansions
SIAM Journal on Scientific Computing, vol. 31, no. 2, pp. 826–848, 2009..